(1) Since

$$\begin{split} \mathbb{E}e^{\eta X} &= \int_0^\infty \mathbb{P}(e^{\eta X} > t)dt = \int_0^1 \mathbb{P}(\Omega)dt + \int_1^\infty \mathbb{P}(e^{\eta X} > t)dt \\ &= 1 + \int_0^\infty \mathbb{P}(e^{\eta X} > e^{\eta s})\eta e^{\eta s}ds = 1 + \eta \int_0^\infty \mathbb{P}(X > s)e^{\eta s}ds \\ &= 1 + \eta \int_0^\infty e^{\eta s}(1 - F(s))ds, \end{split}$$

we get $\int_0^\infty e^{\eta s} (1 - F(s)) ds = \frac{1}{\eta} \left(\mathbb{E} e^{\eta X} - 1 \right).$ (2) (a)

$$\frac{1 - F(x)}{e^{-\lambda x}} = \frac{\left(\frac{b}{x}\right)^a}{e^{-\lambda x}} = b^a \frac{e^{\lambda x}}{x^a} \to \infty$$

as $x \to \text{for any } \lambda > 0$, since the exponential function increases faster than any power function. Thus the distribution is heavy tailed.

(b)

$$\frac{1-F(x)}{e^{-\lambda x}} = \frac{e^{-cx^{\tau}}}{e^{-\lambda x}} = e^{\lambda x - cx^{\tau}} = e^{\lambda x (1-\frac{c}{\lambda}x^{\tau-1})}$$

If $\tau < 1$, then $x^{\tau-1} \to 0$ as $x \to \infty$ and

$$\lim_{x \to \infty} e^{\lambda x (1 - \frac{c}{\lambda} x^{\tau - 1})} = \lim_{x \to \infty} e^{\lambda x} = \infty$$

for all $\lambda > 0$. Hence the Weibull distribution is heavy tailed for $\tau < 1$. If $\tau \ge 1$, then $x^{\tau-1} \ge 1$ for all x > 0 and $-\frac{c}{\lambda}x^{\tau-1} \le -\frac{c}{\lambda}$. Choose $\lambda < c$, so that $1 - \frac{c}{\lambda}x^{\tau-1} \le 1 - \frac{c}{\lambda} < 0$. Then

$$\lim_{x \to \infty} e^{\lambda x (1 - \frac{c}{\lambda} x^{\tau - 1})} \le \lim_{x \to \infty} e^{\lambda x (1 - \frac{c}{\lambda})} = \lim_{x \to \infty} e^{(\lambda - c)x} = 0.$$

Hence the Weibull distribution is light tailed for $\tau \geq 1$.

(3) (a)

$$q_{n} = \binom{m}{n} p^{n} (1-p)^{m-n} = \frac{m!}{(m-n)!n!} p^{n} (1-p)^{m-n}$$

$$= \frac{(m-n+1)m!}{(m-n+1)(m-n)!(n-1)!n} p p^{n-1} \frac{1}{1-p} (1-p)^{m-n+1}$$

$$= \frac{m-n+1}{n} \frac{m!}{(m-n+1)!(n-1)!} \frac{p}{1-p} p^{n-1} (1-p)^{m-n+1}$$

$$= \left(-1 + \frac{m+1}{n}\right) \frac{p}{1-p} \frac{m!}{(m-(n-1))!(n-1)!} p^{n-1} (1-p)^{m-(n-1)}$$

$$= \left(-\frac{p}{1-p} + \frac{(m+1)\frac{p}{1-p}}{n}\right) q_{n-1},$$

so that $a = -\frac{p}{1-p}$ and $b = (m+1)\frac{p}{1-p}$.

(b)

$$\begin{split} q_n &= \binom{v+n-1}{n} p^v (1-p)^n = \frac{(v+n-1)!}{(v+n-1-n)!n!} p^v (1-p)^n \\ &= \frac{(v+n-2)!(v+n-1)}{(v-1)!(n-1)!n} p^v (1-p)^{n-1} \\ &= \frac{v+n-1}{n} (1-p) \frac{(v+n-2)!}{(v-1)!(n-1)!} p^v (1-p)^{n-1} \\ &= \left(1 + \frac{v-1}{n}\right) (1-p) \frac{(v+(n-1)-1)!}{(v+(n-1)-1-(n-1))!(n-1)!} p^v (1-p)^{n-1} \\ &= \left(1 - p + \frac{(v-1)(1-p)}{n}\right) \binom{v+(n-1)-1}{n-1} p^v (1-p)^{n-1} \\ &= \left(1 - p + \frac{(v-1)(1-p)}{n}\right) q_{n-1}, \end{split}$$

so that a = 1 - p and b = (v - 1)(1 - p).

(4) $N(T,\omega) - N(T-,\omega) = 1$ if and only if there is a jump in the path at T, that means $T_n(\omega) = T$ for some $n = 1, 2, \ldots$ Therefore

$$\mathbb{P}(N(T) - N(T-) = 1) = \mathbb{P}(T_n = T \text{ for some } n = 1, 2, \ldots) = \mathbb{P}\left(\bigcup_{n=1}^{\infty} \{T_n = T\}\right)$$
$$= \sum_{n=1}^{\infty} \mathbb{P}(T_n = T).$$

We have $T_n = W_1 + \cdots + W_n$, where W_1, \ldots, W_n are independent and have exponential distribution. Then also T_n has a density function h_n (Exercises -1-) and

$$\mathbb{P}(T_n = T) = \int_{\mathbb{R}} \mathbb{1}_{\{T\}}(x)h_n(x)dx = \int_T^T h_n(x)dx = 0.$$

