

Bayes Teoria, tehtävät-4, ratkaisut

1. Jos y_1, \dots, y_n ovat ehdollisesti riippumattomia ehdolla θ , $p(y|\theta) \sim \mathcal{N}(\theta, v)$ ja $\theta \sim \mathcal{N}(m, w)$ jossa tunnetaan v, w, m silloin a posteriori $\theta|y_1, \dots, y_n \sim \mathcal{N}(m_1, w_1)$

$$w_1 = \left(\frac{1}{w} + \frac{n}{v} \right)^{-1} = \frac{w}{k} \text{ jos ja vain jos}$$
$$k = 1 + n \frac{w}{v}$$

Huomaat muuttuneesta notaatiosta.

2.

$$p(\theta | \text{data}) \sim \exp(-\theta^2/18 + \theta/9)$$

täydentämällä neliön saadaan

$$= \exp\left(-\frac{1}{2} \frac{(\theta - 2)^2}{9}\right) \exp\left(-\frac{2^2}{9}\right)$$

siis $p(\theta | \text{data})$ on gaussinen $\mathcal{N}(2, 9)$.

Sen 90%-HDI on $(2 - 1.644854 \times \sqrt{9}, 2 + 1.644854 \times \sqrt{9}) = (-2.934561, 6.934561)$

3. Olkoon (y_1, \dots, y_n) ovat ehdollisesti riippumattomia ja $\mathcal{N}(\theta, 1)$ jakautuneet ehdolla θ . $\pi(\theta) \propto \mathbf{1}(\theta) > 0$. Tämä priori on epäaito, ei ole todennäköisyysmitta.

Kuitenkin posteriori sattaa olla aito todennäköisyysmitta.

$$p(\theta|y_1, \dots, y_n) \propto \mathbf{1}(\theta > 0) \exp\left(\theta \sum_{i=1}^n y_i - \frac{n}{2} \theta^2\right)$$
$$\propto \mathbf{1}(\theta > 0) \exp\left(-\frac{1}{2/n} (\theta - \bar{y}_n)^2\right)$$

jossa $\bar{y}_n = n^{-1} \sum_{i=1}^n y_i$

Kun $n > 0$, posteriori on ehdollistettu gaussinen jakauma

$$p(\theta|y_1, \dots, y_n) = \frac{\mathbf{1}(\theta > 0)}{1 - \Phi(-\sqrt{n}\bar{y}_n)} \sqrt{\left(\frac{n}{2\pi}\right)} \exp\left(-\frac{1}{2/n} (\theta - \bar{y}_n)^2\right)$$

jossa $\Phi(t) = P(G \leq t)$ on standardi normali gaussisen kertymäfunktio.

$$1 - \Phi(-\sqrt{n}\bar{y}_n) = \Phi(\sqrt{n}\bar{y}_n)$$

4. Olkoon $\varphi = \frac{\theta}{1-\theta}$, $\theta = \frac{\varphi}{1+\varphi}$.

$$\begin{aligned}
\int \mathbf{1}_A(\varphi(\theta))p_\theta(\theta)d\theta &= \int \mathbf{1}_A(\varphi)p_\theta(\theta(\varphi))\frac{d\theta(\varphi)}{d\varphi}d\varphi \\
&= \int \mathbf{1}_A(\varphi)p_\varphi(\varphi)d\varphi \\
p_\varphi(\varphi) &= p_\theta(\theta(\varphi))\frac{d\theta(\varphi)}{d\varphi}
\end{aligned}$$

Koska

$$\frac{d\theta(\varphi)}{d\varphi} = \frac{1}{(1+\varphi)} - \frac{\varphi}{(1+\varphi)^2} = \frac{1}{(1+\varphi)^2}$$

seuraa

$$\begin{aligned}
p_\varphi(\varphi) &= \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+n-s)} \left(\frac{\varphi}{1+\varphi}\right)^{\alpha+s-1} \left(\frac{1}{1+\varphi}\right)^{\beta+n-s-1} \frac{1}{(1+\varphi)^2} \\
&= \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+n-s)} \varphi^{(\alpha+s-1)} \left(\frac{1}{1+\varphi}\right)^{-(\alpha+\beta+n)}
\end{aligned}$$

5.

$$P(X_{n+1} = 1 | X_1 = 1, \dots, X_n = 1) = \frac{\alpha+n}{\alpha+\beta+n}$$

jossa a priori $\theta \sim \text{Beta}(\alpha, \beta)$. Arvot $\alpha = \beta 1$ vastaavat tasaista jakaumaa välissä $[0, 1]$. Kun θ a priori on $[0, 1]$ -tasaisesti jakautunut, seuraa

$$\frac{\alpha+n}{\alpha+\beta+n} = \frac{n+1}{n+2}$$

6.

$$\begin{aligned}
E(y_{n+1}|y_1, \dots, y_n) &= \int_{\mathbb{R}} E(y_{n+1}|\theta)p(\theta|y_1, \dots, y_n)d\theta \\
&= \int_{\mathbb{R}} \theta p(\theta|y_1, \dots, y_n)d\theta = m_1
\end{aligned}$$

$$\begin{aligned}
\text{Var}(y_{n+1}|y_1, \dots, y_n) &= E(\text{Var}(y_{n+1}|\theta)|y_1, \dots, y_n) + \text{Var}(E(y_{n+1}|\theta)|y_1, \dots, y_n) = \\
&= E(1|y_1, \dots, y_n) + \text{Var}(\theta|y_1, \dots, y_n) = 1 + w_1
\end{aligned}$$

7. $y_1, \dots, y_n | \theta$ ehdollisesti riippumattomia jakaumalla $\sim \text{Tasainen}(0, \theta)$. $\pi(\theta) \propto \frac{1}{\theta}, \theta > 0$.

$$p(\theta|y_1, \dots, y_n) \propto \mathbf{1}(\max_i y_i \leq \theta)\theta^{-(n+1)}$$

$$\theta^{-(n+1)} = -\frac{1}{n} \frac{d}{d\theta} \theta^{-n}$$

$$\int_{\max y_i}^{\infty} \theta^{-(n+1)} d\theta = \frac{1}{n} (\max y_i)^{-n}$$

Siis

$$p(\theta|y_1, \dots, y_n) = n \mathbf{1}(\max_i y_i \leq \theta) (\max y_i)^n \theta^{-(n+1)}$$

Kun $n = 6$ ja $\max y_i = 1.96$

$$\begin{aligned} P(\theta > 2.5|y_1, \dots, y_n) &= n(1.96)^n \int_{\max\{2.5, 1.96\}}^{\infty} \theta^{-(n+1)} d\theta \\ &= \left(\frac{1.96}{2.5}\right)^n = 0.2322183 \end{aligned}$$