

Bayes Teoria, ratkaisut-3, (28.01.2011)

1.

$$\begin{aligned} p(\theta_1|Y=0) &= \frac{p(Y=0|\theta_1)p(\theta_1)}{p(Y=0|\theta_1)p(\theta_1) + p(Y=0|\theta_2)p(\theta_2)} \\ &= \frac{5/6 \times 1/2}{5/6 \times 1/2 + 2/3 \times 1/2} = 5/9 \\ p(\theta_2|Y=0) &= 1 - p(\theta_1|Y=0) = 4/9 \\ p(\theta_1|Y=1) &= \frac{p(Y=1|\theta_1)p(\theta_1)}{p(Y=1|\theta_1)p(\theta_1) + p(Y=1|\theta_2)p(\theta_2)} \\ &= \frac{1/6 \times 1/2}{1/6 \times 1/2 + 1/3 \times 1/2} = 1/3 \\ p(\theta_2|Y=1) &= 1 - p(\theta_1|Y=1) = 2/3 \end{aligned}$$

$$\begin{aligned} \frac{p(\theta_1)}{1 - p(\theta_1)} &= \frac{p(\theta_1)}{p(\theta_2)} = \frac{1/2}{1/2} = 1 \\ \frac{p(\theta_1|Y=1)}{1 - p(\theta_1|Y=1)} &= \frac{p(\theta_1|Y=1)}{p(\theta_2|Y=1)} = \frac{1/3}{2/3} = 1/2 \\ \frac{p(\theta_1|Y=0)}{1 - p(\theta_1|Y=0)} &= \frac{p(\theta_1|Y=0)}{p(\theta_2|Y=0)} = \frac{5/9}{4/9} = 5/4 \end{aligned}$$

Kun priori muutetaan $p(\theta_1) = 1 - p(\theta_2) = 5/6$

$$\frac{p(\theta_1)}{1 - p(\theta_1)} = \frac{p(\theta_1)}{p(\theta_2)} = \frac{5/6}{1/6} = 5$$

$$\begin{aligned} p(\theta_1|Y=0) &= \frac{5/6 \times 5/6}{5/6 \times 5/6 + 2/3 \times 1/6} = 25/29 \\ p(\theta_2|Y=0) &= 1 - p(\theta_1|Y=0) = 4/29 \\ p(\theta_1|Y=1) &= \frac{1/6 \times 5/6}{1/6 \times 5/6 + 1/3 \times 1/6} = 5/7 \\ p(\theta_2|Y=1) &= 1 - p(\theta_1|Y=1) = 2/7 \\ \frac{p(\theta_1|Y=0)}{p(\theta_2|Y=0)} &= 25/4 \\ \frac{p(\theta_1|Y=1)}{p(\theta_2|Y=1)} &= 5/2 \end{aligned}$$

2.

$$p(x|\theta) = \begin{cases} 2\theta x \exp(-\theta x^2) & \text{kun } x > 0 \\ 0 & \text{muuten} \end{cases}$$

$$\begin{aligned}
F(x|\theta) &= P(X \leq x|\theta) = \int_0^x 2\theta t \exp(-\theta t^2) dt \\
&= \int_0^x \frac{d}{dt} \exp(-\theta t^2) dt = \left[\exp(-\theta t^2) \right]_0^x = 1 - \exp(-\theta x^2) \\
S(x|\theta) &= P(X > x|\theta) = 1 - F(x|\theta) = \exp(-\theta x^2)
\end{aligned}$$

$$\begin{aligned}
h(x|\theta) &= -\frac{d}{dx} \log S(x|\theta) = 2\theta x \\
&= \lim_{\Delta t \downarrow 0} \frac{1}{\Delta t} P(X \in [t + \Delta t] | X \geq t)
\end{aligned}$$

3.

$$p(x|\theta) = \begin{cases} \theta \exp(-\theta x) & \text{kun } x > 0 \\ 0 & \text{muuten} \end{cases}$$

$$P(X > x|\theta) = \int_x^\infty \theta \exp(-\theta t) dt = \exp(-\theta x)$$

$$P(\sqrt{X} > x|\theta) = P(X > x^2|\theta) = \exp(-\theta x^2)$$

4. Olkoon posteriori jakauma

$$p(\theta | \text{data} = y) = \frac{1}{2} \theta^2 \exp(-\theta), \quad \theta > 0$$

Tämä on gamma jakauma parametreilla (3, 1).

$$\begin{aligned}
E(\theta^n | y) &= \frac{1}{2} \int_0^\infty \theta^{n+2} \exp(-\theta) d\theta = \\
\frac{\Gamma(3+n)}{\Gamma(3)} &= \frac{(2+n)!}{2!}
\end{aligned}$$

jossa kun $\alpha > 0$

$$\Gamma(\alpha) = \int_0^\infty \theta^{\alpha-1} \exp(-\theta) d\theta$$

toteuttaa $\Gamma(1) = 1$, $\Gamma(\alpha + 1) = \alpha \Gamma(\alpha)$ ja siksi $\Gamma(n) = (n-1)!$. Siksi $E(\theta|y) = 3$, $E(\theta^2|y) = 12$, $\text{Var}(\theta|y) = 12 - 3^2 = 3$, siis keskihajonta on $\sqrt{3}$

Voidaan myös laskea jakauman kertymäfunktio. Osittaisintegroinnilla

$$\begin{aligned}
F(t|y) &= P(\theta \leq t|y) = \frac{1}{2} \int_0^t \theta^2 \exp(-\theta) d\theta = \\
&= -\frac{1}{2} \exp(-t)t^2 + \int_0^t \theta \exp(-\theta) d\theta = -\frac{1}{2} \exp(-t)t^2 - \exp(-t)t + \int_0^t \exp(-\theta) d\theta = \\
&= -\frac{1}{2} \exp(-t)t^2 - \exp(-t)t + 1 - \exp(-t) = 1 - \exp(-t)(1 + t + t^2/2)
\end{aligned}$$

90%-symmetrisen luottamus väli on (a, b) jolla $F(a|y) = 0.05$ ja $1 - F(b|y) = 0.05$, Saadaan epälineaarinen yhtälö

$$\exp(-t)(1 + t + t^2/2) = \alpha$$

joka ratkaistaan numeerisesti.

5. 90% HDI: etsitään väli (a, b) jolla

$$p(a|y) = p(b|y) \quad \text{ja} \quad F(b|y) - F(a|y) = \alpha$$

jossa $\alpha = 0.9$. Siis ratkaistaan epälineaarinen yhtälö

$$\begin{aligned} a^2 \exp(-a) &= b^2 \exp(-b) \\ \exp(-a)(1 + a + a^2/2) - \exp(-b)(1 + b + b^2/2) &= \alpha \end{aligned}$$

6. Kun priori on $p(\theta_1) = p(\theta_2) = 1/2$

$$\begin{aligned} p(\tilde{Y} = 1|Y = 1) &= p(\tilde{Y} = 1, \theta_1|Y = 1) + p(\tilde{Y} = 1, \theta_2|Y = 1) = \\ &= p(\tilde{Y} = 1|\theta_1, Y = 1)p(\theta_1|Y = 1) + p(\tilde{Y} = 1|\theta_2, Y = 1)p(\theta_2|Y = 1) \\ &= p(\tilde{Y} = 1|\theta_1)p(\theta_1|Y = 1) + p(\tilde{Y} = 1|\theta_2)p(\theta_2|Y = 1) \\ &= 1/6 \times 1/3 + 1/3 \times 2/3 = 5/18 \end{aligned}$$

Kun priori $p(\theta_1) = 1 - p(\theta_2) = 5/6$

$$p(\tilde{Y} = 1|Y = 1) = 1/6 \times 5/7 + 1/3 \times 2/7 = 3/14$$

7. (Cavendish data)

$$\theta \sim \mathcal{N}(m, w) \quad (\text{priori})$$

$$y_i|\theta \sim \mathcal{N}(\theta, v) \quad 1 \leq i \leq n$$

(ehdollisesti riippumattomia ja samoin jakautuneita),

$$\theta|y_1, \dots, y_n \sim \mathcal{N}(m_1, w_1) \quad (\text{posteriori})$$

$$w_1 = \left(1/w + n/v\right)^{-1}, \quad m_1 = w_1 \left(m/w + n\bar{y}_n/v\right)$$

$$\bar{y}_n = \frac{1}{n} \sum_{i=1}^n y_i = 5.48 \quad n = 23,$$

$$w_1 = (10000/4 + 23 \times 100/4)^{-1} = 4/12300 = 3.252 \times 10^{-4}$$

$$m_1 = 3.252 \times 10^{-4} (5.48 \times 10000/4 + 23 \times 5.48 \times 100/4) =$$

$$\frac{4}{12300} \frac{5.48 \times 12300}{4} = 5.48$$

8. Winbugs koodi:

```
model {  
  theta ~ dnorm( m,w)  
  for(i in 1: n) {  
    y[i] ~ dnorm( theta,v )  
  }  
}  
  
data  
list( n=23, m=5.48, w=0.0004,v=0.04,  
y=c( 5.36,5.29,5.58,5.65,5.57,5.53,5.62,5.29,5.44,5.34,  
5.79,5.1,5.27,5.39,5.42,5.47,5.63,5.34,5.46,5.3,  
5.78,5.68,5.85) )  
inits  
list( theta = 5.48 )
```